

CSCI B609: **“Foundations of Data Science”**

Lecture 5: Dimension Reduction, Separating and Fitting Gaussians

Slides at <http://grigory.us/data-science-class.html>

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Gaussian Annulus Theorem

- Gaussian in d dimensions ($N_d(0^d, 1)$):

$$\Pr[\mathbf{x} = (z_1, \dots, z_d)] = (2\pi)^{-\frac{d}{2}} e^{-\frac{z_1^2 + z_2^2 + \dots + z_d^2}{2}}$$

Nearly all mass in annulus of radius \sqrt{d} and width $O(1)$:

- **Thm.** For any $\beta \leq \sqrt{d}$ all but $3e^{-c\beta^2}$ probability mass satisfies $\sqrt{d} - \beta \leq \|\mathbf{x}\|_2 \leq \sqrt{d} + \beta$ for constant c

Nearest Neighbors and Random Projections

- Given a database A of n points in \mathbb{R}^d
 - Preprocess A into a small data structure D
 - Should answer following queries fast:

Given $\mathbf{q} \in \mathbb{R}^d$ find closest $\mathbf{x} \in A$: $\operatorname{argmin}_{\mathbf{x} \in A} \|\mathbf{q} - \mathbf{x}\|_2$

- Project each $\mathbf{x} \in A$ onto $f(\mathbf{x})$, where $f: \mathbb{R}^d \rightarrow \mathbb{R}^k$
- Pick k vectors $\mathbf{u}_1, \dots, \mathbf{u}_k$ i.i.d: $\mathbf{u}_i \sim N_d(0^d, 1)$
$$f(\mathbf{v}) = (\langle \mathbf{u}_1, \mathbf{v} \rangle, \dots, \langle \mathbf{u}_k, \mathbf{v} \rangle)$$
- Will show that w.h.p. $\|f(\mathbf{v})\|_2 \approx \sqrt{k} \|\mathbf{v}\|_2$

Return: $\operatorname{argmin}_{\mathbf{x} \in A} \|f(\mathbf{q}) - f(\mathbf{x})\|_2 = \operatorname{argmin}_{\mathbf{x} \in A} \|f(\mathbf{q} - \mathbf{x})\|_2 \approx \sqrt{k} \operatorname{argmin}_{\mathbf{x} \in A} \|\mathbf{q} - \mathbf{x}\|_2$

Random Projection Theorem

- Pick k vectors $\mathbf{u}_1, \dots, \mathbf{u}_k$ i.i.d: $\mathbf{u}_i \sim N_d(0^d, 1)$
 $f(\mathbf{v}) = (\langle \mathbf{u}_1, \mathbf{v} \rangle, \dots, \langle \mathbf{u}_k, \mathbf{v} \rangle)$
- Will show that w.h.p. $\|f(\mathbf{v})\|_2 \approx \sqrt{k} \|\mathbf{v}\|_2$

Thm. Fix $\mathbf{v} \in \mathbb{R}^d$ then $\exists c > 0$: for $\epsilon \in (0,1)$:

$$\Pr_{\mathbf{u}_i \sim N_d(0^d, 1)} \left[\left| \|f(\mathbf{v})\|_2 - \sqrt{k} \|\mathbf{v}\|_2 \right| \geq \epsilon \sqrt{k} \|\mathbf{v}\|_2 \right] \leq 3 e^{-ck\epsilon^2}$$

- Scaling: $\|\mathbf{v}\|_2 = 1$
- **Key fact:** $\langle \mathbf{u}_i, \mathbf{v} \rangle = \sum_{j=1}^d \mathbf{u}_{ij} \mathbf{v}_j \sim N(0, \|\mathbf{v}\|_2^2) = N(0,1)$
- Apply “Gaussian Annulus Theorem” with $k = d$

Nearest Neighbors and Random Projections

Thm. Fix $\mathbf{v} \in \mathbb{R}^d$ then $\exists c > 0$: for $\epsilon \in (0,1)$:

$$\Pr_{\mathbf{u}_i \sim N_d(0^d, 1)} \left[\left| \|\mathbf{f}(\mathbf{v})\|_2 - \sqrt{k} \|\mathbf{v}\|_2 \right| \geq \epsilon \sqrt{k} \|\mathbf{v}\|_2 \right] \leq 3 e^{-ck\epsilon^2}$$

Return: $\operatorname{argmin}_{\mathbf{x} \in A} \|\mathbf{f}(\mathbf{q}) - \mathbf{f}(\mathbf{x})\|_2 \approx \sqrt{k} \operatorname{argmin}_{\mathbf{x} \in A} \|\mathbf{q} - \mathbf{x}\|_2$

- Fix and let $\mathbf{v} = \mathbf{q} - \mathbf{x}_i$ for $\mathbf{x}_i \in A$ and let $k = O\left(\frac{\gamma \log n}{\epsilon^2}\right)$
 $(1 \pm \epsilon) \sqrt{k} \|\mathbf{q} - \mathbf{x}_i\|_2 \approx \|\mathbf{f}(\mathbf{q}) - \mathbf{f}(\mathbf{x})\|_2$ (prob. $1 - n^{-\gamma}$)

- Union bound:

For fixed \mathbf{q} distances to A preserved with prob. $1 - n^{-\gamma+1}$

Separating Gaussians

- One-dimensional mixture of Gaussians:
$$p(x) = w_1 p_1(x) + w_2 p_2(x)$$
- E.g. modeling heights of men/women
- **Parameter estimation problem:**
 - Given samples from a mixture of Gaussians
 - **Q:** Estimate means and (co)-variances
- **Sample origin problem:**
 - Given samples from **well-separated** Gaussians
 - **Q:** Did they come from the same Gaussian?

Separating Gaussians

- Gaussian in d dimensions ($N_d(0^d, 1)$):

$$\Pr[\mathbf{x} = (z_1, \dots, z_d)] = (2\pi)^{-\frac{d}{2}} e^{-\frac{z_1^2 + z_2^2 + \dots + z_d^2}{2}}$$

Nearly all mass in annulus of radius \sqrt{d} and width $O(1)$:

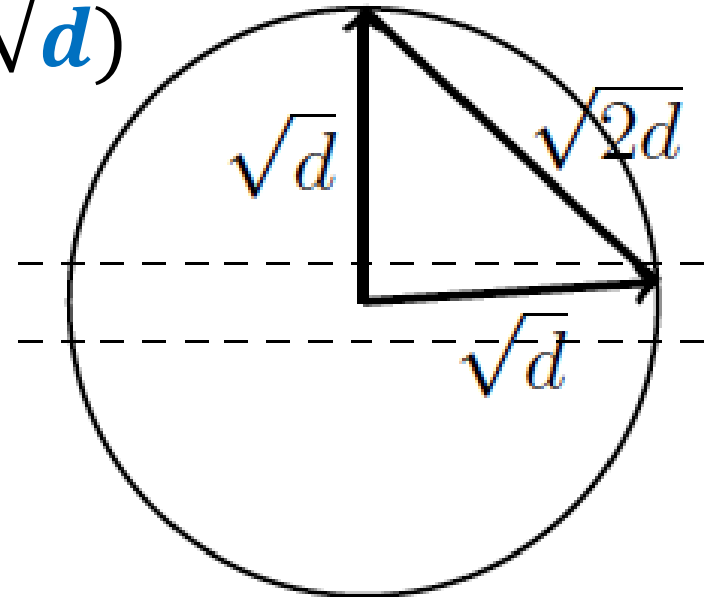
- Almost all mass in a slab $\{\mathbf{x} \mid -c \leq x_1 \leq c\}$ for $c = O(1)$
- Pick $\mathbf{x} \sim$ Gaussian and rotate coordinates to make it x_1
- Pick $\mathbf{y} \sim$ Gaussian, w.h.p. projection of \mathbf{y} on \mathbf{x} is $\in [-c, c]$

$$\|\mathbf{x} - \mathbf{y}\|_2 \approx \sqrt{\|\mathbf{x}\|_2^2 + \|\mathbf{y}\|_2^2}$$

Separating Gaussians

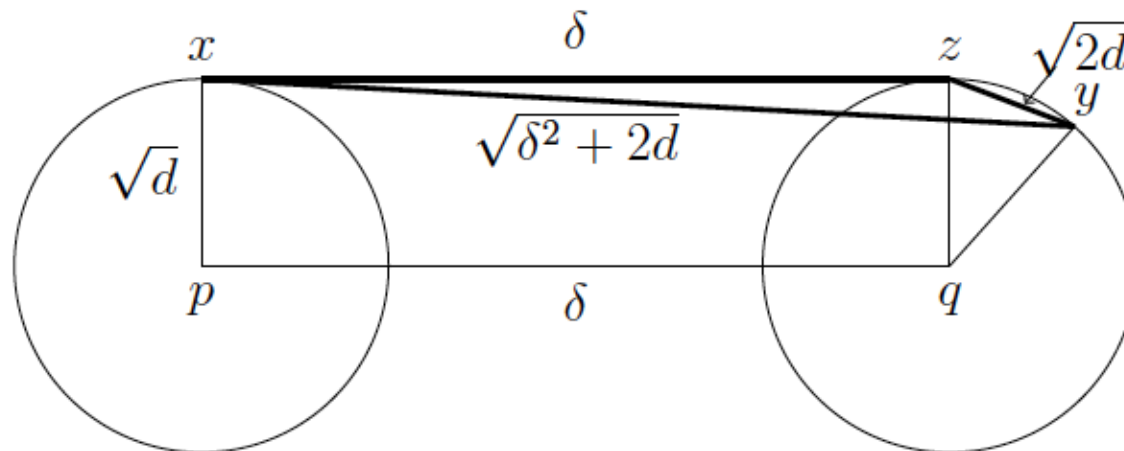
In coordinates:

- $\mathbf{x} = (\sqrt{d} \pm O(1), 0, 0, \dots, 0)$
- $\mathbf{y} = (\pm O(1), \sqrt{d} \pm O(1), 0, \dots, 0)$
- W.h.p: $\|\mathbf{x} - \mathbf{y}\|_2^2 = 2d \pm O(\sqrt{d})$



Separating Gaussians

- Two spherical unit variance Gaussians centered at \mathbf{p}, \mathbf{q}
- $\delta = \|\mathbf{p} - \mathbf{q}\|_2$
- $(\mathbf{x} \sim N(\mathbf{p}, 1), \mathbf{y} \sim N(\mathbf{q}, 1))$
- $\mathbf{x} = (\sqrt{d} \pm O(1), 0, 0, 0, \dots, 0)$
- $\mathbf{y} = (\pm O(1), \delta \pm O(1), \sqrt{d} \pm O(1), 0, \dots, 0)$
- $\|\mathbf{x} - \mathbf{y}\|_2^2 = \delta^2 + 2d \pm O(\sqrt{d})$



Separating Gaussians

- Same Gaussian:

$$\|\mathbf{x} - \mathbf{y}\|_2^2 = 2\mathbf{d} \pm O(\sqrt{\mathbf{d}})$$

- Different Gaussians:

$$\|\mathbf{x} - \mathbf{y}\|_2^2 = \delta^2 + 2\mathbf{d} \pm O(\sqrt{\mathbf{d}})$$

- Separation requires:

$$2\mathbf{d} \pm O(\sqrt{\mathbf{d}}) < \delta^2 + 2\mathbf{d} \pm O(\sqrt{\mathbf{d}})$$

$$O(\sqrt{\mathbf{d}}) < \delta^2$$

$$\omega(\mathbf{d}^{1/4}) < \delta$$

Fitting Spherical Gaussian to Data

- Given samples $\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_n$
- **Q:** What are parameters of best fit $N(\boldsymbol{\mu}, \sigma)$?

$$\begin{aligned}\Pr[\mathbf{x}_i = (z_1, \dots, z_d)] \\ &= (2\pi)^{-\frac{d}{2}} e^{-\frac{(\mu_1 - z_1)^2 + (\mu_2 - z_2)^2 + \dots + (\mu_d - z_d)^2}{2}} \\ &= (2\pi)^{-\frac{d}{2}} e^{-\frac{\|\boldsymbol{\mu} - \mathbf{z}\|_2^2}{2}}\end{aligned}$$

$$\begin{aligned}\Pr[\mathbf{x}_1 = \mathbf{z}_1, \mathbf{x}_2 = \mathbf{z}_2, \dots, \mathbf{x}_n = \mathbf{z}_n] \\ &= (2\pi)^{-\frac{dn}{2}} e^{-\frac{\|\boldsymbol{\mu} - \mathbf{z}_1\|_2^2 + \|\boldsymbol{\mu} - \mathbf{z}_2\|_2^2 + \dots + \|\boldsymbol{\mu} - \mathbf{z}_d\|_2^2}{2\sigma^2}}\end{aligned}$$

Maximum Likelihood Estimator

- PDF: $(2\pi)^{-\frac{dn}{2}} e^{-\frac{\|\mu - z_1\|_2^2 + \|\mu - z_2\|_2^2 + \dots + \|\mu - z_d\|_2^2}{2\sigma^2}}$
- MLE for μ is $\mu = \frac{1}{n} (\mathbf{x}_1 + \mathbf{x}_2 + \dots + \mathbf{x}_n)$
- Take gradient w.r.t μ and make it = 0
- $\nabla_{\mu} \|\mu - \mathbf{x}\|_2^2 = 2(\mu - \mathbf{x})$
 $2(\mu - \mathbf{x}_1) + 2(\mu - \mathbf{x}_2) + \dots + 2(\mu - \mathbf{x}_n) = 0$

MLE for Variance

- $a = \|\boldsymbol{\mu} - \mathbf{x}_1\|_2^2 + \|\boldsymbol{\mu} - \mathbf{x}_2\|_2^2 + \dots + \|\boldsymbol{\mu} - \mathbf{x}_d\|_2^2$
- $\nu = 1/2\sigma^2$
- PDF:
$$\frac{e^{-a\nu}}{\left[\int_{\mathbf{x} \in \mathbb{R}^d} e^{-\nu \|\mathbf{x}\|_2^2} d\mathbf{x} \right]^n}$$
- Log(PDF):
$$-a\nu - n \ln \left[\int_{\mathbf{x} \in \mathbb{R}^d} e^{-\nu \|\mathbf{x}\|_2^2} d\mathbf{x} \right]$$
- Differentiate w.r.t. ν and set derivative = 0

MLE for Variance

- Log(PDF): $-a\boldsymbol{v} - n \ln \left[\int_{\boldsymbol{x} \in \mathbb{R}^d} e^{-\boldsymbol{v} \|\boldsymbol{x}\|_2^2} d\boldsymbol{x} \right]$

- $\frac{d}{d\boldsymbol{v}}$ Log(PDF):

$$-a + n \frac{\int_{\boldsymbol{x} \in \mathbb{R}^d} \|\boldsymbol{x}\|_2^2 e^{-\boldsymbol{v} \|\boldsymbol{x}\|_2^2} d\boldsymbol{x}}{\int_{\boldsymbol{x} \in \mathbb{R}^d} e^{-\boldsymbol{v} \|\boldsymbol{x}\|_2^2} d\boldsymbol{x}}$$

- $y = \|\boldsymbol{v}\boldsymbol{x}\|_2^2$:

$$-a + \frac{n \int_{\boldsymbol{x} \in \mathbb{R}^d} y^2 e^{-y^2} d\boldsymbol{x}}{\boldsymbol{v} \int_{\boldsymbol{x} \in \mathbb{R}^d} e^{-y^2} d\boldsymbol{x}} = -a + \frac{n}{\boldsymbol{v}} \times \frac{d}{2} = 0$$

- $\boldsymbol{v} = \frac{1}{2\sigma^2} \Rightarrow \text{MLE}(\sigma) = \sqrt{\frac{a}{nd}}$ = sample standard deviation